Superposition Codes for Mismatched Decoding

Jonathan Scarlett University of Cambridge jms265@cam.ac.uk Alfonso Martinez Universitat Pompeu Fabra alfonso.martinez@ieee.org Albert Guillén i Fàbregas ICREA & Universitat Pompeu Fabra University of Cambridge guillen@ieee.org

Abstract—An achievable rate is given for discrete memoryless channels with a given (possibly suboptimal) decoding rule. The result is obtained using a refinement of the superposition coding ensemble. The rate is tight with respect to the ensemble average, and can be weakened to the LM rate of Hui and Csiszár-Körner, and to Lapidoth's rate based on parallel codebooks.

I. INTRODUCTION

In this paper, we consider the problem of channel coding over a discrete memoryless channel (DMC) W(y|x), in which the decoder maximizes the symbol-wise product of a given decoding metric q(x, y). If q(x, y) = W(y|x) then the decoder is a maximum-likelihood (ML) decoder; otherwise, the decoder is said to be mismatched [1]–[6]. More precisely, the decoder estimates the message as

$$\hat{m} = \arg\max_{j} \prod_{i=1}^{n} q(x_i^{(j)}, y_i),$$
 (1)

where *n* is the block length, $\boldsymbol{x}^{(j)} = (x_1^{(j)}, \ldots, x_n^{(j)})$ is the *j*-th codeword in the codebook, and $\boldsymbol{y} = (y_1, \ldots, y_n)$ is the received vector.

An error is said to have occurred if the estimated message differs from the transmitted one. A rate R is said to be achievable if, for all $\delta > 0$, there exists a sequence of codebooks with at least $\exp(n(R - \delta))$ codewords and vanishing error probability. The mismatched capacity, defined to be the supremum of all achievable rates, is unknown in general, and most existing work has focused on achievable random-coding rates. Of particular note is the LM rate [3], [5], given by (see Section I-A for notation)

$$I^{\text{LM}}(Q) \stackrel{\triangle}{=} \min_{\substack{\widetilde{P}_{XY} : \widetilde{P}_X = Q, \widetilde{P}_Y = P_Y\\ \mathbb{E}_{\widetilde{P}}[\log q(X,Y)] \ge \mathbb{E}_{F}[\log q(X,Y)]}} I_{\widetilde{P}}(X;Y), \quad (2)$$

where Q is an arbitrary input distribution, and $P_{XY} = Q \times W$. The generalized mutual information (GMI) [6] is defined similarly, with the constraint $\tilde{P}_X = P_X$ removed.

To our knowledge, only two improvements on the LM rate have been reported in the literature. In [4], Csiszár and Narayan show that better achievable rates can be obtained by applying the LM rate to the channel

 $W^{(2)}((y_1, y_2)|(x_1, x_2)) = W(y_1|x_1)W(y_2|x_2)$ with the metric $q^{(2)}((x_1, x_2), (y_1, x_2)) = q(x_1, y_1)q(x_2, y_2)$, and similarly for the k-th order products of W and q. However, as k increases, the required computation becomes prohibitively complex, since the LM rate is non-convex in Q in general. In [1], Lapidoth gives a single-letter improvement on the LM rate using multiple-access coding techniques, in which multiple codebooks are generated in parallel.

In this paper, we obtain a new single-letter improvement on the LM rate using a refinement of superposition coding, the standard version of which is typically used in broadcast scenarios [7]. The results of this paper and the existing singleletter results in the literature can be summarized by the following list of random-coding constructions, in decreasing order of rate (see Sections II-A to II-D for further discussion):

- 1) Refined Superposition Coding (Theorem 1),
- 2) Standard Superposition Coding ((11)-(12)),
- 3) Expurgated Parallel Coding ([1, Theorem 4]),
- 4) Constant-Composition Coding (LM rate [3], [5]),
- 5) i.i.d. Coding (GMI [6]).

A. Notation

The set of all probability distributions on an alphabet \mathcal{A} is denoted by $\mathcal{P}(\mathcal{A})$. The set of all empirical distributions (i.e. types) corresponding to length-*n* sequences on \mathcal{A} is denoted by $\mathcal{P}_n(\mathcal{A})$. The set of all sequences of length *n* with a given type P_X is denoted by $T^n(P_X)$, and similarly for joint types. See [8, Ch. 2] for an introduction to the method of types.

The marginals of a joint distribution $P_{XY}(x, y)$ are denoted by $P_X(x)$ and $P_Y(y)$. Similarly, $P_{Y|X}(y|x)$ denotes the conditional distribution induced by $P_{XY}(x, y)$. We write $P_X = \tilde{P}_X$ to denote element-wise equality between two probability distributions on the same alphabet. Expectation with respect to a distribution $P_X(x)$ is denoted by $\mathbb{E}_P[\cdot]$.

Given a distribution Q(x) and a conditional distribution W(y|x), the joint distribution Q(x)W(y|x) is denoted by $Q \times W$. Information-theoretic quantities with respect to a given distribution (e.g. $P_{XY}(x,y)$) are written with a subscript (e.g. $I_P(X;Y)$). We write $[\alpha]^+ = \max(0,\alpha)$, and denote the indicator function by $\mathbb{1}\{\cdot\}$. Logarithms have base e, and rates are in nats except in the examples, where bits are used.

II. MAIN RESULT

We begin by introducing the refined superposition coding ensemble from which the achievable rate is obtained. We fix a finite alphabet \mathcal{U} , an input distribution Q_{UX} and the rates

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 R_0 and $\{R_{1u}\}_{u \in \mathcal{U}}$. We write $M_0 \stackrel{\triangle}{=} e^{nR_0}$ and $M_{1u} \stackrel{\triangle}{=} e^{nR_{1u}}$. We let $P_{\mathcal{U}}(u)$ be the uniform distribution over the type class $T^n(Q_{U,n})$, where $Q_{U,n} \in \mathcal{P}_n(\mathcal{U})$ is the most probable type under Q_U . For each $u \in \mathcal{U}$, we define

$$n_u \stackrel{\triangle}{=} Q_{U,n}(u)n \tag{3}$$

and let $Q_{u,n_u} \in \mathcal{P}_{n_u}(\mathcal{X})$ be the most probable type under $Q_{X|U=u}$ for sequences of length n_u . We let $P_{\mathbf{X}_u}(\mathbf{x}_u)$ be the uniform distribution over the type class $T^{n_1}(Q_{u,n_u})$. Thus,

$$P_{\boldsymbol{U}}(\boldsymbol{u}) = \frac{1}{|T^n(Q_{\boldsymbol{U},n})|} \mathbb{1}\left\{\boldsymbol{u} \in T^n(Q_{\boldsymbol{U},n})\right\}$$
(4)

$$P_{\boldsymbol{X}_{u}}(\boldsymbol{x}_{u}) = \frac{1}{|T^{n_{u}}(Q_{u,n_{u}})|} \mathbb{1}\left\{\boldsymbol{x}_{u} \in T^{n_{u}}(Q_{u,n_{u}})\right\}.$$
 (5)

We randomly generate the length-*n* auxiliary codewords $\{U^{(i)}\}_{i=1}^{M_0}$ independently according to P_U . For each $i = 1, \ldots, M_0$ and $u \in \mathcal{U}$, we further generate the length- n_u partial codewords $\{X_u^{(i,j_u)}\}_{j_u=1}^{M_{1u}}$ independently according to P_{X_u} . For example, when $\mathcal{U} = \{1, 2\}$ we have

$$\left\{ \left(\boldsymbol{U}^{(i)}, \left\{ \boldsymbol{X}_{1}^{(i,j_{1})} \right\}_{j_{1}=1}^{M_{11}}, \left\{ \boldsymbol{X}_{2}^{(i,j_{2})} \right\}_{j_{2}=1}^{M_{12}} \right) \right\}_{i=1}^{M_{0}} \\ \sim \prod_{i=1}^{M_{0}} \left(P_{\boldsymbol{U}}(\boldsymbol{u}^{(i)}) \prod_{j_{1}=1}^{M_{11}} P_{\boldsymbol{X}_{1}}(\boldsymbol{x}_{1}^{(i,j_{1})}) \prod_{j_{2}=1}^{M_{12}} P_{\boldsymbol{X}_{2}}(\boldsymbol{x}_{2}^{(i,j_{2})}) \right). \quad (6)$$

The codebook on \mathcal{X}^n is indexed as $(m_0, m_{11}, \ldots, m_{1|\mathcal{U}|})$. To transmit a given message, we treat $U^{(m_0)}$ as a timesharing sequence; at the indices where $U^{(m_0)}$ equals u, we transmit the symbols of $X_u^{(m_0,m_{1u})}$. There are $M_0 \prod_u M_{1u}$ codewords, and hence the rate is $R = R_0 + \sum_u Q_{U,n}(u)R_{1u}$. We will see that in the mismatched setting, this ensemble yields higher achievable rates than the standard constantcomposition superposition coding ensemble [7] in which, for all $i = 1, \ldots, M_0$, the codewords $\{X^{(i,j)}\}_{j=1}^{M_1}$ are conditionally independent given $U^{(i)}$.

The main result of this paper is stated in the following theorem, which is proved in Section III. We define the set

$$\mathcal{T}_{0}(P_{UXY}) \stackrel{\triangle}{=} \left\{ \widetilde{P}_{UXY} \in \mathcal{P}(\mathcal{U} \times \mathcal{X} \times \mathcal{Y}) : \widetilde{P}_{UX} = P_{UX}, \\ \widetilde{P}_{Y} = P_{Y}, \mathbb{E}_{\widetilde{P}}[\log q(X,Y)] \ge \mathbb{E}_{P}[\log q(X,Y)] \right\}.$$

$$(7)$$

Theorem 1. For any finite set U and input distribution Q_{UX} , the rate

$$R = R_0 + \sum_{u} Q_U(u) R_{1u}$$
 (8)

is achievable provided that R_0 and $\{R_{1u}\}_{u=1}^{|\mathcal{U}|}$ satisfy

$$R_{1u} \le I^{\rm LM}(Q_{X|U=u}) \tag{9}$$

$$R_{0} \leq \min_{\widetilde{P}_{UXY} \in \mathcal{T}_{0}(Q_{UX} \times W)} I_{\widetilde{P}}(U;Y) + \left[\max_{\mathcal{K} \subseteq \mathcal{U}} \sum_{u \in \mathcal{K}} Q_{U}(u) \Big(I_{\widetilde{P}}(X;Y|U=u) - R_{1u} \Big) \right]^{+}.$$
 (10)

By combining the techniques of [9, Section III-C] and [1, Thm. 3], it can be shown that, subject to minor technical conditions, the achievable rate of Theorem 1 is tight with respect to the ensemble average. That is, Theorem 1 gives the best possible achievable rate for the random coding ensemble under consideration. A similar statement holds for the randomcoding error exponent given in the proof in Section III (cf. (30)). Furthermore, a similar analysis can be applied to the mismatched broadcast channel with degraded message sets, in which a secondary user attempts to recover only the index m_0 .

A. Comparison to Existing Results

The LM rate in (2) is recovered from (8)–(10) by setting $|\mathcal{U}| = 1$ and $R_0 = 0$.

Using standard constant-composition superposition coding [7], a similar (yet simpler) analysis to Section III yields the achievability of all $R = R_0 + R_1$ such that (R_0, R_1) satisfy

$$R_{1} \leq \min_{\substack{\widetilde{P}_{UX} = Q_{UX}, \widetilde{P}_{UY} = P_{UY} \\ \mathbb{E}_{\widetilde{p}}[\log q(X,Y)] \geq \mathbb{E}_{P}[\log q(X,Y)]}} I_{\widetilde{p}}(X;Y|U) \quad (11)$$

$$R_0 + R_1 \le \min_{\substack{\widetilde{P}_{UXY} \in \mathcal{T}_0(Q_{UX} \times W) \\ I_{\widetilde{P}}(U;Y) \le R_0}} I_{\widetilde{P}}(U,X;Y)$$
(12)

This achievability result can be obtained by weakening that of Theorem 1 with the identification $R_1 = \sum_u Q_U(u)R_{1u}$. Roughly speaking, we obtain (11) by summing the $|\mathcal{U}|$ rates in (9) weighted by $Q_U(\cdot)$, identifying the corresponding $|\mathcal{U}|$ joint distributions \widetilde{P}_{XY} in (2) with $\widetilde{P}_{XY|U=u}$, and relaxing the constraints on the metric in (2) to hold on average with respect to Q_U , rather than for each $u \in \mathcal{U}$. Furthermore, we obtain (12) from (10) by replacing the maximum over $\mathcal{K} \subseteq \mathcal{U}$ with the particular choice $\mathcal{K} = \mathcal{U}$, noting that (10) always holds when the minimizing \widetilde{P}_{UXY} satisfies $I_{\widetilde{P}}(U;Y) > R_0$, and using the chain rule for mutual information.

The achievable rate of Lapidoth [1, Thm. 4] was derived by (i) fixing the finite auxiliary alphabets \mathcal{X}_1 and \mathcal{X}_2 , the input distributions $Q_{X_1} \in \mathcal{P}(\mathcal{X}_1)$ and $Q_{X_2} \in \mathcal{P}(\mathcal{X}_2)$, and the function $\phi : \mathcal{X}_1 \times \mathcal{X}_2 \to \mathcal{X}$, (ii) coding for the mismatched multiple-access channel $W'(y|x_1, x_2) \stackrel{\triangle}{=} W(y|\phi(x_1, x_2))$ with the metric $q'(x_1, x_2, y) \stackrel{\triangle}{=} q(\phi(x_1, x_2), y)$ and input distributions Q_{X_1} and Q_{X_2} , and (iii) expurgating all codeword pairs except those of a given joint type. The analysis in [1] yields the achievability of all $R' = R'_1 + R'_2$ such that (R'_1, R'_2) satisfy

$$R'_{1} \leq \min_{\substack{\widetilde{P}_{X_{1}X_{2}} = Q_{X_{1}} \times Q_{X_{2}}\\\widetilde{P}_{X_{2}Y} = P_{X_{2}Y}\\\mathbb{E}_{\tilde{p}}[\log q'] \geq \mathbb{E}_{P}[\log q']}} I_{\tilde{p}}(X_{1};Y|X_{2})$$
(13)

$$R_2' \leq \min_{\substack{\widetilde{P}_{X_1X_2} = Q_{X_1} \times Q_{X_2}\\\widetilde{P}_{X,Y} = P_{X,Y}}} I_{\widetilde{P}}(X_2; Y|X_1) \tag{14}$$

$$\mathbb{E}_{\tilde{P}}[\log q'] \ge \mathbb{E}_{P}[\log q']$$

$$R'_{1} + R'_{2} \leq \min_{\substack{\widetilde{P}_{X_{1}X_{2}} = Q_{X_{1}} \times Q_{X_{2}}, \widetilde{P}_{Y} = P_{Y} \\ \mathbb{E}_{\widetilde{P}}[\log q'] \geq \mathbb{E}_{P}[\log q']} I_{\widetilde{P}}(X_{1};Y) \leq R'_{1}, I_{\widetilde{P}}(X_{2};Y) \leq R'_{2}} I_{\widetilde{P}}(X_{1};Y) \leq R'_{2}, I_{\widetilde{P}}(X_{1};Y) \leq R'_{2}} I_{\widetilde{P}}(X_{1};Y) \leq$$

where $P_{X_1X_2Y} = Q_{X_1} \times Q_{X_2} \times W'$, each minimization is over all $\hat{P}_{X_1X_2Y}$ satisfying the specified constraints, and we write $\mathbb{E}_{\tilde{P}}[\log q']$ as a shorthand for $\mathbb{E}_{\tilde{P}}[\log q'(X_1, X_2, Y)]$.

Proposition 1. After the optimization of \mathcal{U} and Q_{UX} , the achievable rate $R = \max_{R_0,R_1} R_0 + R_1$ described by (11)–(12) is at least as high as the achievable rate $R' = \max_{R'_1,R'_2} R'_1 + R'_2$ described by (13)–(15) with optimized parameters \mathcal{X}_1 , \mathcal{X}_2 , Q_{X_1} , Q_{X_2} and $\phi(x_1, x_2)$.

Proof: We fix the alphabets \mathcal{X}_1 and \mathcal{X}_2 , the distributions $Q_{X_1}(x_1)$ and $Q_{X_2}(x_2)$, and the function $\phi(\cdot, \cdot)$ arbitrarily. We consider the superposition coding parameters $U = X_2$ and

$$Q_{UX}(u,x) = \sum_{x_1} Q_{X_1}(x_1) Q_{X_2}(x_2) \mathbb{1} \{ x = \phi(x_1, x_2) \}.$$
(16)

Since the (U, X) marginal of \tilde{P}_{UXY} is constrained to be equal to Q_{UX} in both (11) and (12), we can equivalently rewrite each optimization as a minimization over $\tilde{P}_{Y|UX}$. The corresponding objectives and constraints depend on $\tilde{P}_{Y|UX}(y|u, x)$ only through $\tilde{P}_{Y|X_1X}(y|x_1, \phi(x_1, x_2))$, which is a conditional distribution on \mathcal{Y} given \mathcal{X}_1 and \mathcal{X}_2 . Thus, the bounds can be weakened by taking the minimizations over all distributions on \mathcal{Y} given \mathcal{X}_1 and \mathcal{X}_2 satisfying similar constraints to (11)–(12). With some simple algebra and by renaming $R_1 = R'_1$ and $R_0 = R'_2$, we obtain the achievability of (R'_1, R'_2) satisfying (13) and (15) with the constraint $I_{\widetilde{P}}(X_1; Y) \leq R'_1$ removed. Repeating these steps with $U = X_1$, $R_1 = R'_2$, $R_0 = R'_1$ and

$$Q_{UX}(u,x) = \sum_{x_2} Q_{X_1}(x_1) Q_{X_2}(x_2) \mathbb{1}\{x = \phi(x_1, x_2)\},$$
(17)

we obtain the achievability of (R'_1, R'_2) satisfying (14) and (15) with the constraint $I_{\widetilde{P}}(X_2; Y) \leq R'_2$ removed.

Finally, we make use of [10, Lemma 1], which states that $R_1 + R_2$ satisfies the inequality in (15) if and only if $R_1 + R_2$ satisfies at least one of two similar inequalities, each of the same form as (15) with one constraint $I_{\tilde{P}}(X_{\nu};Y) \leq R_{\nu}$ ($\nu = 1, 2$) removed. It follows that the union of the above two derived regions contains the region in (13)–(15), and hence the former yields a sum rate at least as high as the latter.

B. Discussion

The benefits of both parallel and superposition coding arise from the dependence among the randomly generated codewords. Under parallel coding without expurgation [1], one generates the codewords $\{X_1^{(i)}\}_{i=1}^{M_1}$ and $\{X_2^{(j)}\}_{j=1}^{M_2}$ independently. One can picture the overall codewords as being arranged in an $M_1 \times M_2$ grid, where the (i, j)-th entry is a deterministic function of $(X_1^{(i)}, X_2^{(j)})$. In this case, every codeword in row *i* depends on $X_1^{(i)}$, and every codeword in column *j* depends on $X_2^{(j)}$. Similarly, under superposition coding, one can arrange the codewords in an $M_0 \times M_1$ grid in which every codeword in the *i*-th row depends on $U^{(i)}$.

The dependence among both rows and columns under parallel coding yields two constraints of the form $I_{\tilde{P}}(X_{\nu};Y) \leq R_{\nu}$ $(\nu = 1, 2)$ in (15), whereas (12) has just one analogous constraint. However, from [10, Lemma 1], at most one of the two constraints affects the minimization for any given rate pair.

Superposition coding allows one to specify a joint composition of (U, X), yielding the constraint $\widetilde{P}_{UX} = Q_{UX}$ in (11)–(12). On the other hand, one cannot specify the joint composition of (X_1, X_2) under parallel coding. However, using the expurgation step of [1], one recovers a codebook in which the joint composition is fixed.

While the ability to choose the joint distribution Q_{UX} in (11)-(12) may appear to give more freedom than the ability to choose Q_{X_1} and Q_{X_2} in (13)-(15), it can be shown that any joint distribution of (X_1, X) (or (X_2, X)) can be induced in the latter setting with the additional freedom in choosing $\phi(\cdot, \cdot)$. This observation suggests that for many (and possibly all) channels and decoding metrics the converse to Proposition 1 holds true, i.e. that the achievable rates of standard superposition and expurgated parallel coding coincide after the full optimization of the parameters. On the other hand, we believe that the local optimization of the former rate over (\mathcal{U}, Q_{UX}) has a lower computational complexity than the local optimization of the latter rate over $(\mathcal{X}_1, \mathcal{X}_2, Q_{X_1}, Q_{X_2}, \phi)$. Since computational complexity generally prohibits the global optimization of the random-coding parameters, the ability to perform such local optimizations is of great interest.

Finally, the advantage of the refined superposition coding ensemble over that of standard superposition coding arises from a stronger dependence among the codewords in a given row. Specifically, unlike the former ensemble, the latter ensemble has codewords in each row which are conditionally independent given given $U^{(i)}$.

C. Example 1: Sum Channel

Given two channels (W_1, W_2) respectively defined on the alphabets $(\mathcal{X}_1, \mathcal{Y}_1)$ and $(\mathcal{X}_2, \mathcal{Y}_2)$, the *sum channel* is defined to be the channel W(y|x) with $|\mathcal{X}| = |\mathcal{X}_1| + |\mathcal{X}_2|$ and $|\mathcal{Y}| = |\mathcal{Y}_1| + |\mathcal{Y}_2|$ such that one of the two subchannels is used on each transmission [11]. One can similarly combine two metrics $q_1(x_1, y_1)$ and $q_2(x_2, y_2)$ to form a *sum metric* q(x, y).

Let (W,q) be the sum channel and metric associated with (W_1,q_1) and (W_2,q_2) , and let \hat{Q}_1 and \hat{Q}_2 be the distributions which maximize the LM rate in (2) on the respective subchannels. We set $\mathcal{U} = \{1,2\}, Q_{X|U=1} = (\hat{Q}_1, \mathbf{0})$ and $Q_{X|U=2} = (\mathbf{0}, \hat{Q}_2)$, where **0** denotes the zero vector. We leave Q_U to be specified.

Combining the constraints $\tilde{P}_{UX} = Q_{UX}$ and $\mathbb{E}_{\tilde{P}}[\log q(X,Y)] \geq \mathbb{E}_{P}[\log q(X,Y)]$ in (7), it is straightforward to show that the minimizing $\tilde{P}_{UXY}(u,x,y)$ in (10) only takes on non-zero values for (u,x,y) such that (i) $u = 1, x \in \mathcal{X}_1$ and $y \in \mathcal{Y}_1$, or (ii) $u = 2, x \in \mathcal{X}_2$ and $y \in \mathcal{Y}_2$, where we assume without loss of generality that the subchannel alphabets are disjoint (i.e. $\mathcal{X}_1 \cap \mathcal{X}_2 = \emptyset$ and $\mathcal{Y}_1 \cap \mathcal{Y}_2 = \emptyset$). It follows that U is a deterministic function of Y under the minimizing \tilde{P}_{UXY} , and hence

$$I_{\widetilde{P}}(U;Y) = H(Q_U) - H_{\widetilde{P}}(U|Y) = H(Q_U).$$
(18)

Therefore, the right-hand side of (10) is lower bounded by $H(Q_U)$. Using (8) and performing a simple optimization of Q_U , it follows that the rate $\log \left(e^{I_1^{\text{LM}}(\hat{Q}_1)} + e^{I_2^{\text{LM}}(\hat{Q}_2)}\right)$ is achievable, where I_{ν}^{LM} is the LM rate for subchannel ν . An analogous result has been proved in the setting of matched decoding using the known formula for channel capacity [11]. It should be noted that the LM rate of (W, q) can be strictly less than $\log \left(e^{I_1^{\text{LM}}(\hat{Q}_1)} + e^{I_2^{\text{LM}}(\hat{Q}_2)}\right)$ even for simple examples (e.g. binary symmetric subchannels).

D. Example 2

We consider the channel and decoding metric described by the entries of the matrices

$$\boldsymbol{W} = \begin{bmatrix} 0.99 & 0.01 & 0 & 0\\ 0.01 & 0.99 & 0 & 0\\ 0.1 & 0.1 & 0.7 & 0.1\\ 0.1 & 0.1 & 0.1 & 0.7 \end{bmatrix}$$
(19)
$$\boldsymbol{q} = \begin{bmatrix} 1 & 0.5 & 0 & 0\\ 0.5 & 1 & 0 & 0\\ 0.05 & 0.15 & 1 & 0.05\\ 0.15 & 0.05 & 0.5 & 1 \end{bmatrix}.$$
(20)

We have intentionally chosen a highly asymmetric channel and metric, since such examples generally yield larger gaps between the various achievable rates. Using an exhaustive search to three decimal places, we find the optimized LM rate to be $R_{\rm LM}^* = 1.111$ bits/use, which is achieved by the input distribution $Q_X^* = (0.403, 0.418, 0, 0.179)$.

The global optimization of (8)–(10) over \mathcal{U} and Q_{UX} is difficult. Setting $|\mathcal{U}| = 2$ and applying local optimization techniques using a number of starting points, we obtained an achievable rate of $R^* = 1.313$ bits/use, with $Q_U =$ $(0.698, 0.302), Q_{X|U=1} = (0.5, 0.5, 0, 0)$ and $Q_{X|U=2} =$ (0, 0, 0.528, 0.472). We denote the corresponding input distribution by $Q_{UX}^{(1)}$.

Applying similar techniques to the standard superposition coding rate in (11)–(12), we obtained an achievable rate of $R_{SC}^* = 1.236$ bits/use, with $Q_U = (0.830, 0.170)$, $Q_{X|U=1} = (0.435, 0.450, 0.115, 0)$ and $Q_{X|U=2} = (0, 0, 0, 0, 1)$. We denote the corresponding input distribution by $Q_{UX}^{(2)}$.

The achievable rates for this example are summarized in Table I, where $Q_{UX}^{(LM)}$ denotes the distribution in which U is deterministic and the X-marginal maximizes the LM rate. While the achievable rate of Theorem 1 coincides with that of (11)–(12) under $Q_{UX}^{(2)}$, the former is significantly higher under $Q_{UX}^{(1)}$. Both types of superposition coding yield a strict improvement over the LM rate.

Our parameters may not be globally optimal, and thus we cannot conclude from this example that Theorem 1 yields a strict improvement over (11)–(12) (and hence over (13)–(15)) after optimizing \mathcal{U} and Q_{UX} . However, since the optimization of the input distribution can be a non-convex problem even for the LM rate, finding the global optimum is computationally infeasible in general. Thus, improvements for fixed $|\mathcal{U}|$ and Q_{UX} are of great interest.

Table I Achievable rates for the mismatched channel in (20)-(21).

Input Distribution	Refined SC	Standard SC
$Q_{UX}^{(1)}$	1.313	1.060
$Q_{UX}^{(2)}$	1.236	1.236
$Q_{UX}^{(m LM)}$	1.111	1.111

III. PROOF OF THEOREM 1

For clarity of exposition, we present the proof in the case that $\mathcal{U} = \{1, 2\}$. The same arguments apply to the general case. We let $\Xi(\boldsymbol{u}, \boldsymbol{x}_1, \boldsymbol{x}_2)$ denote the function for constructing the length-*n* codeword from the auxiliary sequence and partial codewords, and write

$$\boldsymbol{X}^{(i,j_1,j_2)} \stackrel{\triangle}{=} \Xi(\boldsymbol{U}^{(i)}, \boldsymbol{X}_1^{(i,j_1)}, \boldsymbol{X}_2^{(i,j_2)}).$$
(21)

For the remainder of the proof, we let $\boldsymbol{y}_u(\boldsymbol{u})$ denote the subsequence of \boldsymbol{y} corresponding to the indices where \boldsymbol{u} equals \boldsymbol{u} . Furthermore, we define $q^n(\boldsymbol{x}, \boldsymbol{y}) \stackrel{\triangle}{=} \prod_{i=1}^n q(x_i, y_i)$ and $W^n(\boldsymbol{y}|\boldsymbol{x}) \stackrel{\triangle}{=} \prod_{i=1}^n W(y_i|x_i)$.

Upon receiving y, the decoder forms the estimate

$$(\hat{m}_{0}, \hat{m}_{1}, \hat{m}_{2}) = \arg\max_{(i,j_{1},j_{2})} q^{n} (\boldsymbol{x}^{(i,j_{1},j_{2})}, \boldsymbol{y})$$

$$= \arg\max_{(i,j_{1},j_{2})} q^{n_{1}} (\boldsymbol{x}^{(i,j_{1})}_{1}, \boldsymbol{y}_{1}(\boldsymbol{u}^{(i)})) q^{n_{2}} (\boldsymbol{x}^{(i,j_{2})}_{2}, \boldsymbol{y}_{2}(\boldsymbol{u}^{(i)})),$$

$$(23)$$

where the objective in (23) follows by separating the indices where u = 1 from those where u = 2. By writing the objective in this form, it is easily seen that for any given *i*, the pair (j_1, j_2) with the highest metric is the one for which j_1 maximizes $q^{n_1}(\boldsymbol{x}_1^{(i,j_1)}, \boldsymbol{y}_1(\boldsymbol{u}^{(i)}))$ and j_2 maximizes $q^{n_2}(\boldsymbol{x}_2^{(i,j_2)}, \boldsymbol{y}_2(\boldsymbol{u}^{(i)}))$. Therefore, we can split the error event into three (not necessarily disjoint) events:

(Type 0) $\hat{m}_0 \neq m_0$

(*Type 1*)
$$\hat{m}_0 = m_0$$
 and $\hat{m}_{11} \neq m_{11}$

(Type 2) $\hat{m}_0 = m_0$ and $\hat{m}_{12} \neq m_{12}$

We denote the corresponding random-coding error probabilities by $\overline{p}_{e,0}$, $\overline{p}_{e,1}$ and $\overline{p}_{e,2}$ respectively.

From the construction of the random-coding ensemble, the type-1 error probability $\overline{p}_{e,1}$ is precisely that of the single-user constant-composition ensemble with rate R_{11} , length $n_1 = nQ_U(1)$, and input distribution $Q_{X|U=1}$. A similar statement holds for the type-2 error probability $\overline{p}_{e,2}$, and the analysis for these error events is identical to the derivation of the LM rate [3], [5], yielding the rate conditions in (9). For the remainder of this section, we analyze the type-0 event.

We assume without loss of generality that $(m_0, m_1, m_2) = (1, 1, 1)$. We let U and X be the codewords corresponding to (1, 1, 1), and let \overline{U} , \overline{X}_1 and \overline{X}_2 be the codewords corresponding to an arbitrary message with $m_0 \neq 1$. For the index *i* corresponding to \overline{U} , we write $\overline{X}_1^{(j_1)}$, $\overline{X}_2^{(j_2)}$ and $\overline{X}^{(j_1,j_2)}$ in place of $X_1^{(i,j_1)}$, $X_2^{(i,j_2)}$ and $X^{(i,j_1,j_2)}$ respectively. Thus, from (21), we have $\overline{X}_1^{(j_1,j_2)} = \Xi(\overline{U}, \overline{X}_1^{(j_1)}, \overline{X}_2^{(j_2)})$.

The error probability for the type-0 event satisfies

$$\overline{p}_{e,0} \leq \mathbb{P}\left[\bigcup_{i \neq 1} \bigcup_{j_1, j_2} \left\{ \frac{q^n(\boldsymbol{X}^{(i,j_1,j_2)}, \boldsymbol{Y})}{q^n(\boldsymbol{X}, \boldsymbol{Y})} \geq 1 \right\} \right], \quad (24)$$

where $(Y|X = x) \sim W^n(\cdot|x)$. Writing the probability as an expectation given (U, X, Y) and applying the truncated union bound, we obtain

$$\overline{p}_{e,0} \leq \mathbb{E} \left[\min \left\{ 1, (M_0 - 1) \right. \\ \left. \times \mathbb{E} \left[\mathbb{P} \left[\bigcup_{j_1, j_2} \left\{ \frac{q^n(\overline{\boldsymbol{X}}^{(j_1, j_2)}, \boldsymbol{Y})}{q^n(\boldsymbol{X}, \boldsymbol{Y})} \geq 1 \right\} \left| \overline{\boldsymbol{U}} \right] \left| \boldsymbol{U}, \boldsymbol{X}, \boldsymbol{Y} \right] \right\} \right],$$
(25)

where we have written the probability of the union over j_1 and j_2 as an expectation given \overline{U} .

Let the joint types of (U, X, Y) and $(\overline{U}, \overline{X}^{(j_1, j_2)}, Y)$ be denoted by P_{UXY} and \widetilde{P}_{UXY} respectively. We claim that

$$\frac{q^n(\overline{\boldsymbol{X}}^{(j_1,j_2)},\boldsymbol{Y})}{q^n(\boldsymbol{X},\boldsymbol{Y})} \ge 1$$
(26)

if and only if

$$\widetilde{P}_{UXY} \in \mathcal{T}_{0,n}(P_{UXY}) \stackrel{\triangle}{=} \mathcal{T}_0(P_{UXY}) \cap \mathcal{P}_n(\mathcal{U} \times \mathcal{X} \times \mathcal{Y}),$$
(27)

where $\mathcal{T}_0(P_{UXY})$ is defined in (7). The constraint $\tilde{P}_{UX} = P_{UX}$ follows from the construction of the random coding ensemble, $\tilde{P}_Y = P_Y$ follows since $(\boldsymbol{U}, \boldsymbol{X}, \boldsymbol{Y})$ and $(\overline{\boldsymbol{U}}, \overline{\boldsymbol{X}}^{(j_1, j_2)}, \boldsymbol{Y})$ share the same \boldsymbol{Y} sequence, and $\mathbb{E}_{\tilde{P}}[\log q(X, Y)] \geq \mathbb{E}_P[\log q(X, Y)]$ coincides with the condition in (26). Thus, expanding (25) in terms of types yields

$$\overline{p}_{e,0} \leq \sum_{P_{UXY}} \mathbb{P}\Big[(\boldsymbol{U}, \boldsymbol{X}, \boldsymbol{Y}) \in T^{n}(P_{UXY}) \Big] \min \left\{ 1, \\
(M_{0} - 1) \sum_{\widetilde{P}_{UXY} \in \mathcal{T}_{0,n}(P_{UXY})} \mathbb{P}\Big[(\overline{\boldsymbol{U}}, \boldsymbol{y}) \in T^{n}(\widetilde{P}_{UY}) \Big] \\
\times \mathbb{P}\Big[\bigcup_{j_{1}, j_{2}} \Big\{ (\overline{\boldsymbol{u}}, \overline{\boldsymbol{X}}^{(j_{1}, j_{2})}, \boldsymbol{y}) \in T^{n}(\widetilde{P}_{UXY}) \Big\} \Big] \Big\}, \quad (28)$$

where we write $(\overline{\boldsymbol{u}}, \boldsymbol{y})$ to denote an arbitrary pair such that $\boldsymbol{y} \in T^n(P_Y)$ and $(\overline{\boldsymbol{u}}, \boldsymbol{y}) \in T^n(\widetilde{P}_{UY})$. The dependence of these sequences on P_Y and \widetilde{P}_{UY} is kept implicit.

Using a similar argument to the discussion following (23), we observe that $(\overline{u}, \overline{X}^{(j_1, j_2)}, y) \in T^n(\widetilde{P}_{UXY})$ if and only if $(\overline{X}_u^{(j_u)}, y_u(\overline{u})) \in T^{n_u}(\widetilde{P}_{XY|U=u})$ for u = 1, 2. Thus, for any subset \mathcal{K} of \mathcal{U} , we can upper bound the probability that $(\overline{u}, \overline{X}^{(j_1, j_2)}, y) \in T^n(\widetilde{P}_{UXY})$ by the probability that $(\overline{X}_u^{(j_u)}, y_u(\overline{u})) \in T^{n_u}(\widetilde{P}_{XY|U=u})$ for all $u \in \mathcal{K}$. By further

upper bounding via the union bound, we obtain

$$\mathbb{P}\left[\bigcup_{j_{1},j_{2}}\left\{\left(\overline{\boldsymbol{u}},\overline{\boldsymbol{X}}^{(j_{1},j_{2})},\boldsymbol{y}\right)\in T^{n}(\widetilde{P}_{UXY})\right\}\right]\leq\min\left\{1,\\ \min_{u=1,2}M_{1u}\mathbb{P}\left[\left(\overline{\boldsymbol{X}}_{u},\boldsymbol{y}_{u}(\overline{\boldsymbol{u}})\right)\in T^{n_{u}}(\widetilde{P}_{XY|U=u})\right],\\ M_{11}M_{12}\mathbb{P}\left[\bigcap_{u=1,2}\left\{\left(\overline{\boldsymbol{X}}_{u},\boldsymbol{y}_{u}(\overline{\boldsymbol{u}})\right)\in T^{n_{u}}(\widetilde{P}_{XY|U=u})\right\}\right]\right\},$$
(29)

where the four terms in the minimization correspond to the four subsets of $\{1, 2\}$.

Substituting (29) into (28), applying the property of types in [8, Ex. 2.3(b)] multiple times, and using the fact that the number of joint types is polynomial in n, we obtain

$$\lim_{n \to \infty} -\frac{1}{n} \log \overline{p}_{e,0}$$

$$\geq \min_{P_{UXY}} \min_{\widetilde{P}_{UXY} \in T_0(P_{UXY})} D(P_{UXY} || Q_{UX} \times W) + \left[I_{\widetilde{P}}(U;Y) + \left[\max_{\mathcal{K} \subseteq \mathcal{U}} \sum_{u \in \mathcal{K}} Q_U(u) \left(I_{\widetilde{P}}(X;Y|U=u) - R_{1u} \right) \right]^+ - R_0 \right]^+,$$
(30)

where the minimization over P_{UXY} is subject to $P_{UX} = Q_{UX}$. Taking $P_{UXY} \rightarrow Q_{UX} \times W$, we obtain that the righthand side of (30) is positive whenever (10) holds with strict inequality, thus concluding the proof.

One can show that (30) holds with equality by following the steps of [9, Section III-C]. Thus, the analysis presented in this section is exponentially tight for the given ensemble.

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